



High-Dimensional Sparse Clustering via Iterative Semidefinite Programming Relaxed K-Means

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Sparse Clustering Problem

Goal: Partition unlabeled observations $\mathbf{X}_1, \dots, \mathbf{X}_n \in \mathbb{R}^p$ into clusters G_1, \dots, G_K ($n \ll p$, known K)

Sparse Gaussian Mixture Model:

- Each observation \mathbf{X}_i belongs to a cluster G_k with mean $\boldsymbol{\mu}_k \in \mathbb{R}^p$:

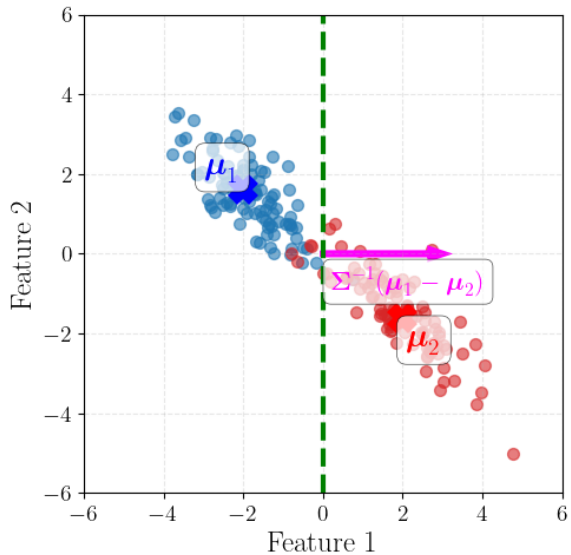
$$\mathbf{X}_i \sim N(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}) \quad \text{if} \quad \mathbf{X}_i \in G_k$$

- Fisher LDA boundary between G_k and G_ℓ : $\boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_k - \boldsymbol{\mu}_\ell)$
- Sparsity assumption: there exist **signal features** and **noise features**

$$S_0 := \bigcup_{k \neq \ell} \text{supp}(\boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_k - \boldsymbol{\mu}_\ell)), \quad \{1, \dots, p\} \setminus S_0$$



Sparsity assumption

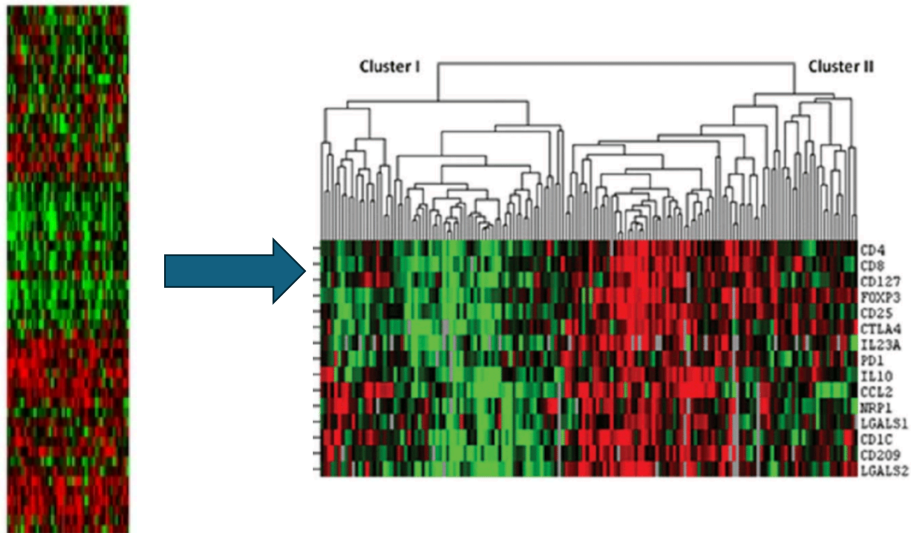


$$\mu_1 - \mu_2 = \begin{pmatrix} -4 \\ 3.2 \end{pmatrix} \text{ is not sparse.}$$

$$\begin{aligned} \Sigma^{-1}(\mu_1 - \mu_2) &= \begin{pmatrix} 1 & -0.8 \\ -0.8 & 1 \end{pmatrix}^{-1} \begin{pmatrix} -4 \\ 3.2 \end{pmatrix} \\ &= \begin{pmatrix} -4 \\ 0 \end{pmatrix} \text{ is sparse.} \end{aligned}$$



Example: disease subtype discovery from gene expression





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	Our method	IF-PCA	SKM	SAS
Leukemia	0.93	0.84	0.79	0.87

- Leukemia dataset: $n = 45$, $p = 3871$, $K = 2$
- Cluster data with labels hidden, evaluate accuracy with true labels
- Baselines: IF-PCA (feature selection \rightarrow clustering), SKM and SAS (iteratively alternate over feature selection and clustering)



Our approach

- **SDP K-means (Peng and Wei, 2005)**
 - Avoids explicit cluster-center estimation
 - Minimax optimal in fixed-dimension, non-sparse regimes (**Chen et al., 2021**)
- **Our contributions**
 - Motivating theory: Extend the analysis of Chen et al. 2021 to sparse setting to study the role of sparsity and feature selection on SDP K-means
 - Extend SDP K-means into an iterative, sparsity-aware algorithm for known covariance setting
 - Extend our algorithm into unknown covariance setting, using the high-dimensional precision matrix estimation tool

K-means (NP-hard)

$$\begin{aligned} \min_{G_1, \dots, G_K} \quad & \sum_{k=1}^K \sum_{i \in G_k} \|\mathbf{X}_i - \bar{\mathbf{X}}_{G_k}\|_2^2 \\ \text{s.t.} \quad & G_1 \cup \dots \cup G_K = \{1, \dots, n\} \\ & G_k \cap G_\ell = \emptyset \text{ for } k \neq \ell \end{aligned}$$

SDP Relaxed K-means

$$\begin{aligned} \max_{\mathbf{Z} \in \mathbb{R}^{n \times n}} \quad & \langle \mathbf{X}^\top \mathbf{X}, \mathbf{Z} \rangle_F \\ \text{s.t.} \quad & \mathbf{Z} = \mathbf{Z}^\top, \mathbf{Z} \succeq 0, \mathbf{Z} \geq 0, \\ & \text{tr}(\mathbf{Z}) = K, \mathbf{Z} \mathbf{1}_n = \mathbf{1}_n \end{aligned}$$

Equivalent matrix form (NP-hard)

$$\begin{aligned} \max_{\mathbf{H} \in \{0,1\}^{n \times K}} \quad & \langle \mathbf{X}^\top \mathbf{X}, \mathbf{H} \mathbf{B} \mathbf{H}^\top \rangle_F, \text{ s.t. } \mathbf{H} \mathbf{1}_K = \mathbf{1}_n \\ \mathbf{X} \in \mathbb{R}^{p \times n} \quad & \text{(data matrix), } \mathbf{B} := (\text{diag}(\mathbf{1}_n^\top \mathbf{H}))^{-1} \end{aligned}$$

$\mathbf{Z} = \mathbf{H} \mathbf{B} \mathbf{H}^\top$ satisfies:

symmetric, PSD, nonnegative, trace K , row sum 1

**True cluster in combinatorial problem is
block-diagonal**

Let G_1^*, \dots, G_K^* be true clusters.

The corresponding decision variable is:

$$\mathbf{Z}^* = \mathbf{H}^* \mathbf{B}^* \mathbf{H}^{*\top} =$$

$$\begin{bmatrix} \frac{1}{|G_1^*|} \mathbf{1}_{|G_1^*| \times |G_1^*|} & 0 & \cdots & 0 \\ 0 & \frac{1}{|G_2^*|} \mathbf{1}_{|G_2^*| \times |G_2^*|} & \cdots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & 0 & \cdots & \frac{1}{|G_K^*|} \mathbf{1}_{|G_K^*| \times |G_K^*|} \end{bmatrix}$$

In practice:

Run spectral clustering on $\hat{\mathbf{Z}}$

**Optimal point of
SDP problem:**

$\hat{\mathbf{Z}}$ = Continuous matrix s.t.

Symmetric, PSD, nonnegative, trace K , row sum 1

Theory (Chen et al. 2021):

If $\min_{1 \leq k \neq \ell \leq K} \|\boldsymbol{\mu}_\ell - \boldsymbol{\mu}_k\|_2^2$
is large enough, with high
probability, $\hat{\mathbf{Z}}$ is exactly \mathbf{Z}^*

For any feature subset $S \subseteq [p]$,

Let $\hat{\mathbf{Z}}(S)$ denote the solution of the SDP corresponding to S :

$$\max_{\mathbf{Z} \in \mathbb{R}^{n \times n}} \langle \mathbf{X}_{S,\cdot}^\top, \mathbf{X}_{S,\cdot}, \mathbf{Z} \rangle_F \quad \text{s.t.} \quad \mathbf{Z} = \mathbf{Z}^\top, \mathbf{Z} \succeq 0, \text{tr}(\mathbf{Z}) = K, \mathbf{Z}\mathbf{1}_n = \mathbf{1}_n, \mathbf{Z} \geq 0.$$

Collection of strong signal feature subsets:

$$\mathcal{S} := \left\{ S \subset [p] : \min_{1 \leq k \neq \ell \leq K} \|(\boldsymbol{\mu}_\ell - \boldsymbol{\mu}_k)_{S \cap S_0}\|_2^2 \gtrsim \left(\log n + \frac{|S| \log p}{n} + \sqrt{\frac{|S| \log p}{n}} \right) \right\}$$

Theorem (Uniform recovery of restricted SDPs)

Assume $\boldsymbol{\Sigma} = \mathbf{I}_p$. Then for any distribution instance in our model,

$$\mathbb{P}\left(\hat{\mathbf{Z}}(S) = \mathbf{Z}^*, \forall S \in \mathcal{S}\right) \gtrsim 1 - \frac{K}{n}.$$

Theorem (Tightness of the required separation)

Assume $\Sigma = \mathbf{I}_p$. There exists a distribution instance in our model such that

1. $\min_{1 \leq k \neq \ell \leq K} \|(\boldsymbol{\mu}_\ell^* - \boldsymbol{\mu}_k^*)_{S_0}\|_2^2 = C \log n$, where C is a constant,
2. For any clustering method f , $\mathbb{P}(f(\mathbf{X}_{S_0, \cdot}) \neq \{G_1^*, \dots, G_K^*\}) \gtrsim 1 - \frac{1}{n}$

If we consider uniform recovery problem restricted to moderately sized subsets satisfying $|S| \lesssim (n \log n) / \log p$:

$$\log n + \frac{|S| \log p}{n} + \sqrt{\frac{|S| \log p}{n}} \asymp \log n.$$

Then SDP K-means is optimal in terms of required separation



Intuition from the theorem

- Simple scenario: $K = 2$, $(\boldsymbol{\mu}_1^* - \boldsymbol{\mu}_2^*)_{S_0} = \mu_0 \mathbf{1}_{|S_0|}$, and $|S| \lesssim (n \log n) / \log p$
- Exact recovery for all S is possible if and only if

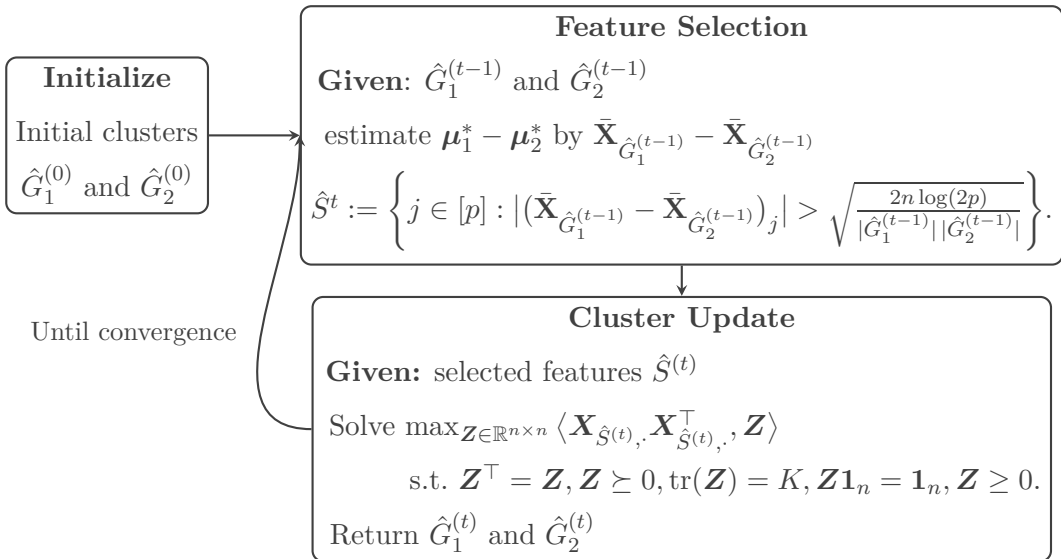
$$|S \cap S_0| \mu_0^2 \gtrsim \log n + \frac{|S| \log p}{n} + \sqrt{\frac{|S| \log p}{n}}.$$

Insights:

1. Features should be chosen based on $\min_{1 \leq k \neq \ell \leq K} \|(\boldsymbol{\mu}_\ell^* - \boldsymbol{\mu}_k^*)_{S \cap S_0}\|_2^2$
2. Mild under- or over-selection is acceptable; severe misselection is harmful
3. Once the algorithm reaches a high-signal subset S , reliable clustering follows



Our Iterative Method under identity covariance





Extension to Unknown Common Covariance

Assumptions:

- Each row of Σ^{-1} has at most J nonzero off-diagonal entries
- There exists a small subset of relevant features $S_0 \subseteq \{1, \dots, p\}$ with $|S_0| \ll p$, such that

$$S_0 := \bigcup_{k \neq \ell} \text{supp}(\Sigma^{-1}(\mu_k - \mu_\ell)) \subset [p].$$

Feature selection step: Replace $\mu_1 - \mu_2$ estimation with

$$\Sigma^{-1}(\mu_1 - \mu_2)$$



Unknown Covariance: Clustering Step

SDP K-Means for general Σ (Zhuang et al 2023), without sparsity:

$$\begin{aligned} \max_{\mathbf{Z} \in \mathbb{R}^{n \times n}} \quad & \left\langle (\Sigma^{-1} \mathbf{X})^\top \Sigma (\Sigma^{-1} \mathbf{X}), \mathbf{Z} \right\rangle \\ \text{s.t.} \quad & \mathbf{Z}^\top = \mathbf{Z}, \mathbf{Z} \succeq 0, \text{tr}(\mathbf{Z}) = K, \mathbf{Z} \mathbf{1}_n = \mathbf{1}_n, \mathbf{Z} \geq 0. \end{aligned}$$



Unknown Covariance: Clustering Step

Given selected features $\hat{S}^{(t)}$, we solve SDP with sub-matrices:

$$\begin{aligned} \max_{\mathbf{Z} \in \mathbb{R}^{n \times n}} \quad & \left\langle (\boldsymbol{\Sigma}^{-1} \mathbf{X})_{\hat{S}^{(t)}, \cdot}^{\top}, \boldsymbol{\Sigma}_{\hat{S}^{(t)}, \hat{S}^{(t)}} (\boldsymbol{\Sigma}^{-1} \mathbf{X})_{\hat{S}^{(t)}, \cdot}, \mathbf{Z} \right\rangle \\ \text{s.t.} \quad & \mathbf{Z}^{\top} = \mathbf{Z}, \mathbf{Z} \succeq 0, \text{tr}(\mathbf{Z}) = K, \mathbf{Z} \mathbf{1}_n = \mathbf{1}_n, \mathbf{Z} \geq 0 \end{aligned}$$



What We Need to Estimate

For both the feature selection and clustering steps, the key quantity required for the extension is

$$\Sigma^{-1}(\mu_1 - \mu_2) \quad \text{and} \quad \Sigma^{-1}X.$$

We *do not* need to explicitly estimate the full precision matrix Σ^{-1} .



Our approach: nodewise regression

We adapt the **Innovated Scalable Efficient Estimation (ISEE; Fan and Lv 2016)**

Idea: Partition the feature indices $[p]$ into disjoint subsets A_1, A_2, \dots, A_m . For each subset A , estimate

$$\Sigma^{-1}\mathbf{X} =$$

$(\Sigma^{-1}\mathbf{X})_{A,\cdot}$
\dots
\dots
\dots

$$\Sigma^{-1}\boldsymbol{\mu}_k =$$

$(\Sigma^{-1}\boldsymbol{\mu}_k)_{A,\cdot}$
\dots
\dots
\dots

by nodewise regression.



Nodewise regression

By multivariate Gaussian assumption $\mathbf{X}_i \stackrel{iid}{\sim} N(\boldsymbol{\mu}_k, \boldsymbol{\Sigma})$,

$$\underbrace{(\mathbf{X}_i)_A}_{\text{response}} = \underbrace{(\boldsymbol{\mu}_k)_A + \boldsymbol{\Omega}_{A,A}^{-1} \boldsymbol{\Omega}_{A,A^c} (\boldsymbol{\mu}_k)_{A^c}}_{\text{intercept}} - \underbrace{\boldsymbol{\Omega}_{A,A}^{-1} \boldsymbol{\Omega}_{A,A^c} \mathbf{X}_{A^c,i}}_{\text{slope}} + \underbrace{\mathbf{E}_{A,i}}_{\text{residual}}, \text{ where } \mathbf{E}_{A,i} \sim N(\mathbf{0}, \boldsymbol{\Omega}_{A,A}^{-1}).$$

$$\boldsymbol{\mu}_k = \begin{array}{|c|} \hline \text{red} \\ \hline \text{blue} \\ \hline \end{array} \begin{array}{l} (\boldsymbol{\mu}_k)_A \\ (\boldsymbol{\mu}_k)_{A^c} \end{array} \quad \boldsymbol{\Omega} = \begin{array}{|c|c|} \hline \text{red} & \text{green } \boldsymbol{\Omega}_{A,A^c} \\ \hline \text{red } \boldsymbol{\Omega}_{A,A} & \dots \\ \hline \end{array} \quad \mathbf{X}_i = \begin{array}{|c|} \hline \text{red} \\ \hline \text{blue} \\ \hline \end{array} \begin{array}{l} (\mathbf{X}_i)_A \\ (\mathbf{X}_i)_{A^c} \end{array}$$



Nodewise regression

- $(\Sigma^{-1} \mu_k)_A = \underbrace{\Omega_{A,A}}_{\text{Cov(residual)}} \underbrace{\left((\mu_k)_A + \Omega_{A,A}^{-1} \Omega_{A,A^c} (\mu_k)_{A^c} \right)}_{\text{intercept}}$
- $(\Sigma^{-1} X_i)_A = (\Sigma^{-1} \mu_k)_A + \underbrace{\Omega_{A,A}}_{\text{Cov(residual)}} \underbrace{E_{A,i}}_{\text{residual}}$

$$\mu_k = \begin{array}{|c|} \hline \text{red} \\ \hline \text{blue} \\ \hline \end{array} \begin{array}{l} (\mu_k)_A \\ (\mu_k)_{A^c} \end{array}$$

$$\Omega = \begin{array}{|c|c|c|} \hline \text{red} & \text{green} & \\ \hline \text{red} & \dots & \\ \hline & & \dots \\ \hline \end{array}$$

$\Omega_{A,A}$ (points to the red block in the first row, first column)

$$X_i = \begin{array}{|c|} \hline \text{red} \\ \hline \text{blue} \\ \hline \end{array} \begin{array}{l} (X_i)_A \\ (X_i)_{A^c} \end{array}$$



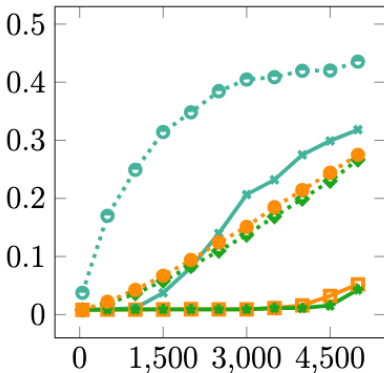
Our Iterative Method under sparse precision matrix

1. **Initialize:** Obtain initial cluster assignments $\hat{G}_1^0, \hat{G}_2^0 \subset [n]$.
2. **Iterate for** $t = 0, 1, 2, \dots$, until convergence:
 - 2.1 **ISEE subroutine:** Given $\hat{G}_1^{(t)}$ and $\hat{G}_2^{(t)}$, estimate $\Sigma^{-1}(\mu_1^* - \mu_2^*)$, $\Sigma^{-1}\mathbf{X}$,
 - 2.2 **Feature selection:** Let \hat{S}^{t+1} be features where estimated $\Sigma^{-1}(\mu_1^* - \mu_2^*)$ vector has large magnitude.

Threshold defined by ℓ_2 convergence rate of ISEE
 - 2.3 **Cluster update:** Run SDP-relaxed K-means on the selected features $\tilde{\mathbf{X}}_{\hat{S}^{t+1}}$, $\Sigma_{\hat{S}^{t+1}, \hat{S}^{t+1}}$ to estimate new clusters $\hat{G}_1^{t+1}, \hat{G}_2^{t+1}$.



Simulations: identity covariance

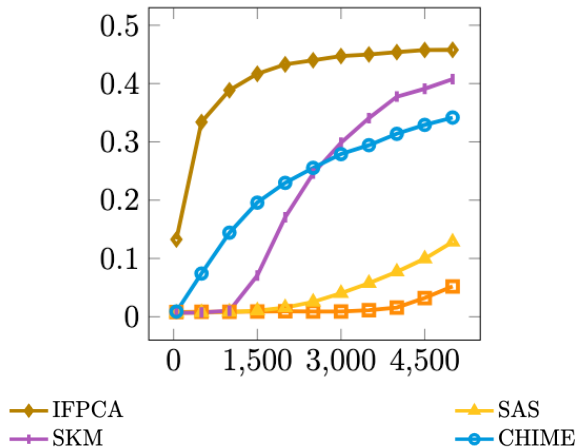


- (1) Spectral clustering (non-sparse)
- (2) Hierarchical clustering (non-sparse)
- (3) SDP K -means (non-sparse)
- Algorithm 2 initialized by (1)
- Algorithm 2 initialized by (2)
- Algorithm 2 initialized by (3)

- X-axis: Dimension p increases while $|S_0| = 10$ and $\|\mu_1 - \mu_2\|_2^2 = 5^2$ are fixed
- Y-axis: mis-clustering rate
- Our method improves upon the non-sparsity-aware baseline.
- Dependence on initial clustering is mild.



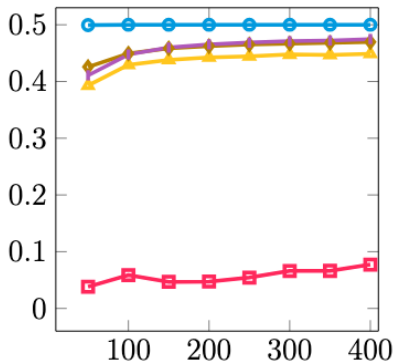
Simulations: identity covariance



- X-axis: Dimension p increases while $|S_0| = 10$ and $\|\mu_1 - \mu_2\|_2^2 = 5^2$ are fixed
- Y-axis: mis-clustering rate
- Our method outperforms existing two-step and iterative methods



Simulations: effectiveness of ISEE and threshold

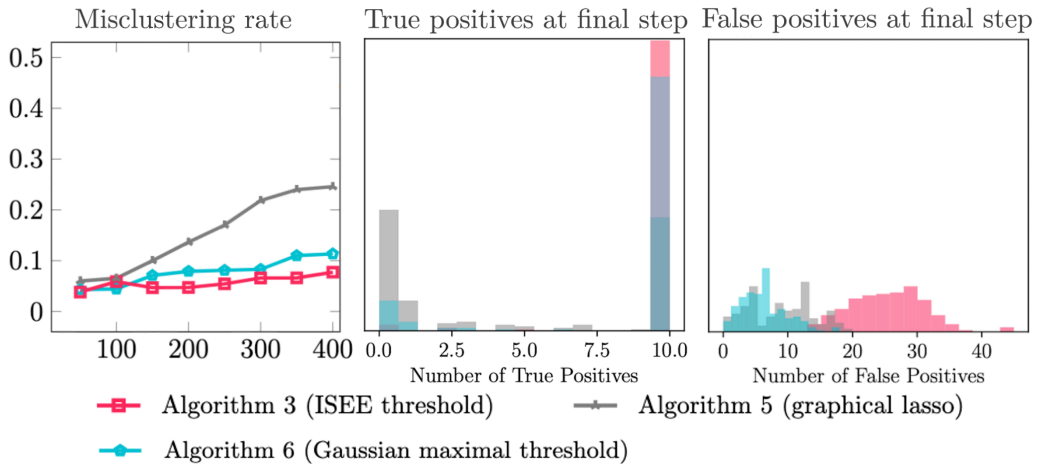


Algorithm 3 SAS IFPCA SKM CHIME

- Σ^{-1} : chain graph with correlation 0.45
- X-axis: Dimension p increases while $|S_0| = 10$, $\|\Sigma^{-1}(\mu_1 - \mu_2)\|_2^2 = 4^2$ fixed
- Y-axis: mis-clustering rate
- Our method outperforms existing two-step and iterative methods



Simulations: sparse precision matrix





Real Data Analysis

	Our method	IFPCA	SKM	SAS
Leukemia	0.93	0.84	0.79	0.87
MNIST	0.94	0.61	0.57	0.56

- Leukemia: $n = 45$, $p = 3871$
- MNIST: $n = 1000$, $p = 784$, digits 1 and 7
- Evaluation metric: clustering accuracy



Conclusion

Sparsity-aware iterative clustering combining convex relaxation, feature selection, and high-dimensional precision estimation.

- Theory-guided:
 - SDP K-means achieves simultaneous exact recovery on feature subsets with sufficient signal; signal requirement is optimal under mild assumptions.
 - Mild under- or over-selection is acceptable; aggressive misselection is harmful.
- Algorithm highlights:
 - Alternates between feature selection (via estimated Fisher LDA) and clustering (SDP K-means).
 - Nodewise regression (ISEE) avoids full precision matrix estimation.



Future Extensions

- Key insight: Mild under- or over-selection is tolerable, but aggressive misselection is harmful
- Current limitation: Past clustering and feature selection results are not explicitly utilized
- Proposed improvements: Introduce explicit exploration and memory
 - Use Thompson sampling for randomized feature selection
 - Update Beta distributions to retain memory of past results
 - Random draws from the Beta distributions encourage exploration



Preliminary simulation for Thompson sampling approach

